Institute for Applied Mathematics SS 2023 Prof. Dr. Anton Bovier, Florian Kreten



Sheet 4, "Stochastic Analysis"

To hand in until May 01, 15:00

For all following exercises, let $\{X_i\}_{i \in \mathbb{N}}$ be an i.i.d sequence, where X_1 has distribution F. Define the extremal process $M_n := \max_{i \leq n} X_i$.

Problem 1 (Examples of extremal limits 3 Pt)

Find out to which domain of attraction the extremal extremal processes M_n of the following distributions belong. Then find suitable sequences $\{a_n\}_{n\in\mathbb{N}} > 0$ and $\{b_n\}_{n\in\mathbb{N}}$, such that $\mathbb{P}[M_n \leq a_n x + b_n]$ converges to a non-degenerate limit.

- 1. F(x) = x, the uniform distribution on (0, 1).
- 2. $F(x) = 1 Kx^{-\alpha}$ for $\alpha > 0, K > 0$, the Pareto distribution on $[K^{1/\alpha}, \infty)$.
- 3. $F(x) = 1 e^{-x}$, the exponential distribution with rate 1 on $[0, \infty)$.

Hint: For the exponential distribution, find a suitable non-linear scaling s(x, n) of M_n , then use a Taylor-Expansion to get a_n and b_n .

Problem 2 (Extremal limit of the normal distribution, 3 Pt)

Proof that if $F = \Phi$ is the standard normal distribution, then

$$\mathbb{P}[M_n \le a_n x + b_n] \to \exp(-e^{-x}), \text{ where}$$

$$a_n = (2\log n)^{-1/2}, \quad b_n = (2\log n)^{1/2} - \frac{1}{2}(2\log n)^{-1/2}(\log\log n + \log 4\pi).$$

Hint: You may use without proof that $1 - \Phi(x) \sim \frac{\phi(x)}{x}$ as $x \to \infty$. As a starting point, find a suitable non-linear scaling.

Problem 3 (Distributions with degenerate extremal limit 1, 2 Pt)

Define $x_F := \sup\{x : F(x) < 1\}$. Assume that $x_F < \infty$ and that $\mathbb{P}[X_1 = x_F] > 0$. Show that for any sequence u_n , such that $\mathbb{P}[M_n \le u_n] \to \rho$, it must be that $\rho \in \{0, 1\}$.

This statement can be extended to the following Theorem, that gives a precise requirement on the tail of the distribution F. There exists a relatively short proof based on Lemma 3.4 in the lecture notes. It is left as a challenging bonus exercise.

Theorem. For any $\rho \in (0,1)$, there exists a sequence $\{u_n\}_{n \in \mathbb{N}}$ with a degenerate limit $\lim_{n \to \infty} P[M_n \leq u_n] \to \rho$, if and only if

$$\lim_{x \to X_F} \frac{1 - F(x)}{1 - F(x_-)} = 1,$$

where $F(x_{-}) := \mathbb{P}[X < x]$, or equivalently if and only if

$$\lim_{x \to X_F} \frac{F(x) - F(x_-)}{1 - F(x_-)} = 0.$$

Problem 4 (Distributions with degenerate extremal limit 2, 2 Pt)

Use the previous statement to conclude that the following distributions do not have a nondegenerate extremal limit:

1) The Poisson distribution with parameter $\lambda > 0$:

$$\mathbb{P}[X=k] = e^{-\lambda} \frac{\lambda^k}{k!} \text{ for } k \in \mathbb{N}_0.$$

2) The Geometric distribution with acceptance probability $p \in (0, 1)$:

$$\mathbb{P}[X=k] = (1-p)^{(k-1)} \cdot p \text{ for } k \in \mathbb{N}_0.$$